

Introduction

- ▶ Obtaining reliable uncertainties for ocean heat content (OHC) is crucial for tracking climate change ($\sim 91\%$ of excess heat in the Earth system is stored in the ocean)
- ▶ Estimating OHC and its uncertainties from Argo temperature profiles presents exciting statistical and computational challenges from **nonstationarity** and **large dataset size**
- ▶ Due to data availability, we split the analysis into two vertical sections (top: 15–975 m and bottom: 975–1850 m), which produces a conservative uncertainty estimate:

$$\text{Var}(\text{OHC}_{\text{total}}|\text{data}) \leq (\sqrt{\text{Var}(\text{OHC}_{\text{top}}|\text{data})} + \sqrt{\text{Var}(\text{OHC}_{\text{bot}}|\text{data})})^2$$

- ▶ Here we improve the uncertainty estimates by extending the locally stationary spatio-temporal interpolation method (Kuusela and Stein, 2018) to model the **vertical dependence**

Mapping methodology

- ▶ **Quantity of interest**

$$\text{OHC}(t) = \rho_0 c_{p,0} \iiint T(x, y, z, t) dx dy dz = \rho_0 c_{p,0} \iiint \underbrace{\left(\int_{z_{\min}}^{z_{\max}} T(x, y, z, t) dz \right)}_{:= \tilde{T}(x, y, t)} dx dy,$$

where $T(x, y, z, t)$ is the potential temperature at space-time location (x, y, z, t)

We map the mean-subtracted temperature residual $\tilde{R}(x, y, t)$ in two vertical sections jointly as realizations from a bivariate Gaussian process (GP)

- ▶ **Bivariate locally stationary GP regression**

Let $s = (x, y, t)$ and $i = \text{year}, j = \text{observation}$. Then we model the temperature residuals as:

$$\begin{bmatrix} \tilde{R}_{\text{top}} \\ \tilde{R}_{\text{bot}} \end{bmatrix}_{i,j} = f_i \left(\begin{bmatrix} s_{\text{top}} \\ s_{\text{bot}} \end{bmatrix}_{i,j} \right) + \begin{bmatrix} \epsilon_{\text{top}} \\ \epsilon_{\text{bot}} \end{bmatrix}_{i,j}$$

$$f_i \stackrel{\text{iid}}{\sim} \text{GP} \left(\begin{bmatrix} 0 \\ 0 \end{bmatrix}, \mathcal{K}(s_1, s_2; \theta) \right), \begin{bmatrix} \epsilon_{\text{top}} \\ \epsilon_{\text{bot}} \end{bmatrix}_{i,j} \stackrel{\text{iid}}{\sim} \mathcal{N} \left(\begin{bmatrix} 0 \\ 0 \end{bmatrix}, \Sigma(\theta) \right)$$

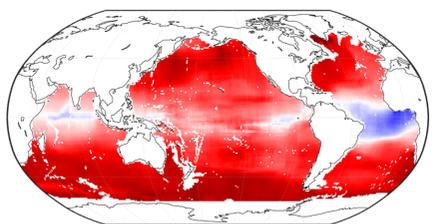
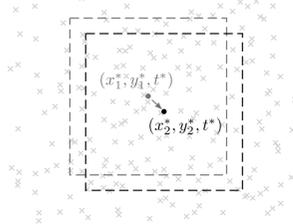
with cross-covariance function (Kleiber and Nychka, 2012)

$$\mathcal{K}_{\text{top,bot}}(s_1, s_2; \theta) = \beta \frac{\delta_{\text{top}} \delta_{\text{bot}}}{\sqrt{|\Theta_{\text{top,bot}}|}} \exp \left(-\sqrt{(s_1 - s_2)^T \Theta_{\text{top,bot}}^{-1} (s_1 - s_2)} \right)$$

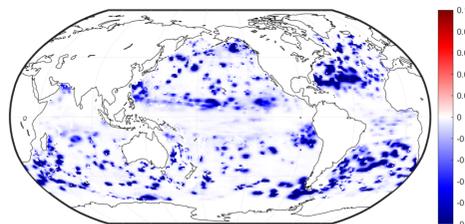
$\Rightarrow \beta$ is an estimate of the dependence between vertical sections!

Parameter estimation and kriging

- ▶ We focus on data in a small neighborhood around a spatio-temporal grid point $s^* = (x^*, y^*, t^*)$ and regard each year as an i.i.d. replicate from the above stationary Gaussian process within the neighborhood
- ▶ We compute the MLE of the covariance parameters and the kriging prediction at s^* for overlapping moving windows (Haas, 1990, 1995), which allows for spatially-varying parameter estimates and predictions



Estimated β (cross-correlation) from 2004–2022 Argo data

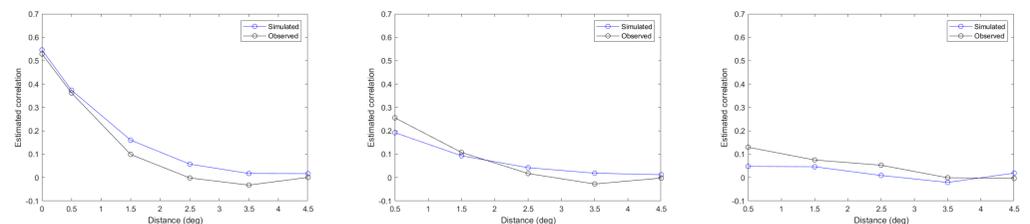


Relative kriging variance difference between bivariate & univariate models (975–1850 m)

\Rightarrow Since the sparser bottom layer is able to borrow strength from the top, we see that the bivariate model provides a **10%+ reduction** in predictive uncertainty

Uncertainty quantification with local conditional simulations

- ▶ Getting uncertainties for $\text{OHC}(t) = \rho_0 c_{p,0} \iint \tilde{T}(x, y, t) dx dy$ requires taking into account the predictive covariance between $\tilde{T}(x_1, y_1, t_1)$ and $\tilde{T}(x_2, y_2, t_2)$
- ▶ We do this by generating ensembles of conditional simulation realizations from the predictive distribution $p(\tilde{T}(\cdot, \cdot, \cdot) | \text{data})$ by convolving a bivariate white noise field with the square roots of local predictive covariance matrices

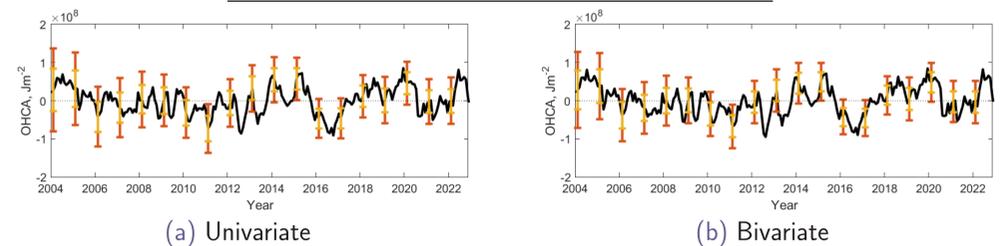


(a) $\Delta t = 0$ months (b) $\Delta t = 1$ month (c) $\Delta t = 2$ months
Cross-validated cross-correlation at varying spatial/temporal distances (x-axis, panels a,b,c) show conditional simulations (blue) capture observed dependence (black) well between vertical sections

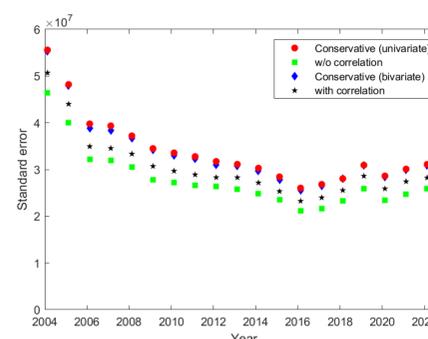
Uncertainty quantification results

- ▶ By producing bivariate conditional simulations, we can now include the cross-dependence term effect in the uncertainty estimate:
 $\text{Var}(\text{OHC}_{\text{total}}|\text{data}) = \text{Var}(\text{OHC}_{\text{top}}|\text{data}) + \text{Var}(\text{OHC}_{\text{bot}}|\text{data}) + 2\text{Cov}(\text{OHC}_{\text{top}}, \text{OHC}_{\text{bot}}|\text{data})$
- ▶ Below we provide uncertainty estimates for the total OHC anomaly (15–1850 m) for Feb of each year from 2004–2022

Global ocean heat content anomalies



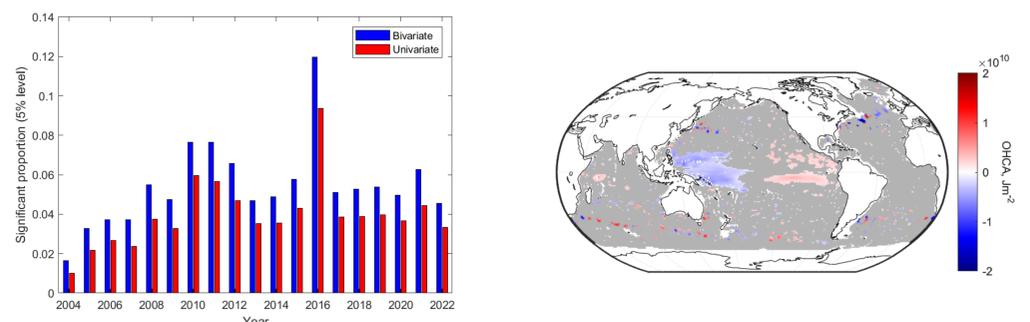
Global OHC anomalies with 68% (orange) and 95% (red) uncertainties



Global OHC anomaly standard errors

The bivariate uncertainties with vertical dependence (black) are **15% smaller** than the univariate conservative (red)

Regional ocean heat content anomalies



Proportion of grid points with significant OHC anomalies at the 5% level

Regional OHC anomalies in Feb 2016; only trends that are significant at 5% level are shown

Ongoing work

- ▶ This framework can also be used to improve uncertainties for OHC trends, ocean heat uptake (OHU), and other quantities of interest to ocean/climate scientists
- ▶ Our univariate estimates with uncertainties were recently featured in the 2023 State of the Climate:

